WORKING ACROSS TIME ZONES: EXPORTERS AND THE GENDER WAGE GAP*

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Abstract

This study argues that there is a systematic difference in the gender wage gap (GWG) between exporting firms and non-exporters. Exporters may require greater commitment from their employees, such as working particular hours to communicate with partners in different time zones or travelling at short notice, and may therefore disproportionately reward employee flexibility. If women are less flexible, or perceived as such, exporters will exhibit a higher GWG than non-exporters. This hypothesis is examined using matched employer-employee data from the Norwegian manufacturing sector for 1996-2010. The results suggest a firm's entry into exporting increases the GWG by about 3 percentage points for college educated workers. A lower overlap in business hours between the Norwegian exporter and its foreign markets and a greater need need for interactions with foreign buyers are associated with a higher GWG.

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1 Introduction

The link between globalization and income inequality has received a lot of attention in economic research as well as in the popular press. However, the debate has tended to ignore one important dimension of inequality, namely, the gender wage gap (GWG). It is a well-documented fact that women earn less than men, even in relatively equal societies such as the Scandinavian countries. This is true after controlling for observable worker characteristics, hours worked and occupation, and especially so in the private sector.¹ In this paper, we set out to explore one channel through which globalization may affect gender inequality—the relationship between exporting and the GWG.

We argue that by virtue of being exposed to higher competition and doing business with partners located in faraway countries, exporters require greater commitment to work and greater flexibility of their employees. For instance, working for an exporting firm may require working particular hours and taking late night phone calls to communicate with customers in different time zones and may involve international travel arranged at short notice. The employees may be expected to be available around the clock seven days a week in case of unexpected problems arising on the production line or shipments being delayed. If women are less flexible, or are perceived as such, exporting firms will exhibit a higher GWG than non-exporters, particularly among skilled workers.

While the relationship between internationalization, new work models and gender has received scant attention in the economics literature to date, our argument is supported by sociological studies on the impact of internationalization and competition on work organization and work practices. Kvande (2009) studies firms with a global reach, and states that technology, which awards flexibility by increasing employee autonomy over time and place of work, leads to increased expectations with respect to availability via phone and e-mail and a need to adjust working hours to international business partners. Formal contracts regulating working hours appear to be replaced by moral obligations and time norms that demand total commitment.² Blair-Loy and Jacobs (2003) interview stock brokers, and conclude that "the majority report that the pace of work has increased in recent years, in part due to the increased competition and new opportunities wrought by new technologies and globalization. [...] These work demands likely also serve as a barrier to women's entering or staying in the occupation." Our study is also motivated by employer surveys and experiments. An employer survey from Iceland reveals that women are perceived as more family oriented than men, less committed to their work, and less

¹See Blau and Kahn (2000), Barth et al. (2017) and Goldin (2014) for overviews.

²An employee of a global company reported that "[d]epending on what country he is working with, he has to adjust his time for fathering. When he was working with Malaysia he lost time with his children in the mornings, because he had to be online with their [Malaysian] time. When he is working with Austria the time with the kids in the afternoon and evening has to be adjusted to the working hours in Austria." When asked whether he sees his job as boundless he agrees: "Yes, and for me it is also very concrete, or physically, because wherever I go I always take my PC and my mobile phone. [...] This weekend, for example, when we were on our Sunday walk in the woods, someone phoned, and then I had to take out my PC and find out something. And with the time difference in Malaysia, which is six actually seven hours before us . . . it is obvious, isn't it, it's boundless." (Kvande, 2009 pp. 68-69).

reliable than men.³

To investigate the link between the GWG and exporting, we exploit a matched employer-employee data set covering the universe of joint stock firms in the Norwegian manufacturing sector and their full-time employees between 1996 and 2010.⁴ The richness of the data allows us to control for observable, time-varying characteristics of the workers and firms that might otherwise confound the effects of interest. The panel structure allows us to take account of unobservable features of the worker, firm, or the worker-firm match that otherwise would bias our results. Norway provides a suitable setting for our study. It has a flexible labor market, as reflected in being ranked in the 9th place in terms of labor market efficiency in the most recent Global Competitiveness Report prepared by the World Economic Forum. It has high quality data which offer the possibility of merging the customs exports figures with the linked employer-employee information. The data are not subject to censoring of either trade activities or recorded wages, as is often the case in comparable data sets. Finally, Norway is a small, open economy with a high trade intensity.

We estimate a Mincerian wage regression controlling for a host of worker and firm characteristics. When we control for observable worker characteristics, but do not account for unobservable heterogeneity, the GWG appears to be smaller in exporting firms than in non-exporting firms. Hence, working for an exporting firm is associated with a smaller observed GWG. However, once we also control for unobservable heterogeneity related to the worker-firm matches, this result is reversed. We find that exporting firms exhibit a higher GWG than non-exporters with the difference between the two types of firms reaching about 3 percentage points. As expected, this effect is present only among college graduates. In other words, college educated women earn higher wages at exporting firms than at non-exporters, as compared to their male colleagues, but they are underpaid given their unobservable characteristics. This result is robust to controlling for selection of firms into exporting with firm-year fixed effects. Firm-year fixed effect also absorb any firm-level characteristics (observed or unobserved, time-varying or not) that could confound our results, such as productivity shocks or changes to the management of the firm.

Exporters vary by size as well as complexity and geography of their operations. These are the features that are likely to affect a firm's need for commitment and flexibility from its workers. Therefore, they should matter for the GWG. We investigate the role of heterogeneity among exporters by analysing the effect of export intensity, the number of export destinations and the number of product varieties on the GWG. Again we condition on both job-spell and firm-year fixed effects. We find that an increase in the number of export destinations, the share of output exported or the number of exported varieties are all associated with a higher GWG. To account for heterogeneity in geography, we calculate the overlap in business hours between the Norwegian firm and its export markets.⁵ Based on our hypothesis, we expect the need for client

 $^{^3}$ See Gislason (2007) and http://www.slatergordon.co.uk/media-centre/news/2014/08/slater-gordon-highlights-maternity-discrimination.

⁴The data set covers about 90 percent of manufacturing output in Norway in 2004.

⁵For firms exporting more than one product to one market, we take the average overlap in business hours

communications outside of normal business hours to matter and to depend on the geography of the firm's export markets. In line with our expectations, a smaller overlap in the business hours between the Norwegian exporter and its foreign markets is associated with a higher GWG. Since there is a negative correlation between the distance to Norway and the attitudes towards females, we check whether our results are robust to controlling for the extent of gender equality in the export market. We find that this does not seem to matter, suggesting that our results capture aspects related to physical distance and different time zones rather than the ability of females to function in the destination country. Finally, we show that the overlap in business hours between the Norwegian firm and its export markets has a greater impact on the GWG among workers under 45 years of age. Women in this age group are likely to have small children which limits their flexibility with respect to working hours. Thus this additional evidence is also in line with our hypothesis.

An alternative explanation for our findings is based on Becker's theory of taste-based discrimination, which predicts that more profitable firms are better positioned to engage in costly discrimination (Becker (1957)). We conduct robustness checks to rule out that this is driving our results. Namely, we show that our findings are robust to allowing the GWG to vary with firm profitability as well as other correlates of profitability, such as firm size and multinational status.

In summary, our findings suggest that firms with a larger exposure to foreign markets and thus a greater need for communications with partners located in different time zones, exhibit a higher GWG. Our findings are reminiscent of Goldin (2014) who argues that professions requiring workers to work particular hours exhibit a higher GWG. However, while Goldin (2014) draws a distinctions between different professions (lawyers versus pharmacists) and sectors (finance versus high-tech) we draw a distinction between different types of firms based on the nature of their activities. Our results add a new angle to Goldin's hypothesis by highlighting that firms requiring more flexibility and commitment from their workers exhibit a greater wage differential between men and women.

Our paper speaks to two main strands of the existing literature in economics. The first strand is the large literature on exporters and the exporter wage premium. For instance, Bernard et al. (1995) show that exporters pay higher wages than non-exporters in the US and that this wage premium goes to both production and non-production workers. More recently, the availability of matched employer-employee data has allowed researchers to examine whether the exporter wage premium is robust to controlling for various worker and firm characteristics. Among others, Schank et al. (2007), using matched employer-employee data for the German manufacturing sector, show that the wage differential becomes smaller but does not vanish once these controls are included. A study by Irarrazabal et al. (2013) on Norwegian manufacturing firms finds that exporters' wage and TFP premia fall substantially once one controls for observable and unobservable worker characteristics, suggesting that exporters attract more able workers than

across the firm's varieties (i.e., country-product combinations).

non-exporters. In contrast, Krishna et al. (2014) find no exporter premium using Brazilian data. A related literature explores the link between trade liberalization in the export markets and export sector wages, though its conclusions are mixed (Revenga (1997), Trefler (2004), Pavcnik et al. (2004), Goldberg and Pavcnik (2005), Goldberg and Pavcnik (2007) and Amiti and Davis (2012)).

The second strand is the large literature on the GWG. Blau and Kahn (2000) and OECD (2012) offer extensive reviews of this literature while recent contributions include Barth et al. (2017), Card et al. (2016) and Hirsch et al. (2010). As mentioned above, of particular interest is a recent paper by Goldin (2014) who argues that gender wage gap is largely due to employers disproportionately rewarding individuals for laboring long hours and working particular hours. Most closely related to our paper are studies specifically on the impact of globalization on the GWG. Black and Brainerd (2004) test Becker's model of discrimination by comparing the difference in residual wage gaps between concentrated and competitive industries that were exposed to comparable increases in import competition. They find that the gender wage gap narrowed more quickly in the originally more concentrated industries and interpret this result as supporting the classic prediction of Becker's theory. Juhn et al. (2014) examine how trade liberalization affects gender inequality using Mexican firm-level data. Trade liberalization induces the most productive firms to start exporting and adopt a new technology that is more modern and requires less physical strength. This improves women's labor market outcomes in the blue-collar tasks, while leaving them unchanged in the white-collar tasks. The authors use the introduction of NAFTA to test the model. They find that liberalization is associated with entry of new firms and adoption of a more modern technology at the same time as firms hire relatively more bluecollar females and increase their share of the wage bill. Finally, a theoretical contribution by Ben Yahmed (2013), based on a Melitz-type framework, predicts that trade integration widens the GWG among skilled workers.⁶ To the best of our knowledge, ours is the first study directly investigating the relationship between one key feature of globalization, namely exporting, and the gender wage gap, and doing so using matched employer-employee data.

The rest of the paper is structured as follows. The next section describes the data and the construction of our data set as well as variables. Section 3 presents the empirical model. Section 4 starts by presenting baseline results and then focuses on heterogeneity among exporters and robustness checks. Section 5 concludes.

⁶In her setting, men and women have the same distribution of ability but differ in their commitment levels. Firms engage in statistical discrimination which creates the GWG. Firm may choose to invest in high technology by incurring an additional cost. Complementarities between commitment/ability and technology induce exporters (which are the ones choosing to adopt the high technology) to hire more able workers and have a higher GWG among skilled workers.

2 Data and Measurement

2.1 Data

We have access to a rich employer-employee data set that is based on administrative registers from Statistics Norway and covers the entire working population of Norway for the period 1996-2010 with detailed information on a set of demographic characteristics like education and family.

The main source of employment and wage data for the period 1996 to 2010 is the Employee Register (AT) which holds annual records of worked hours and earned wages at the individual level. Statistics Norway links this register to the tax office database (LTO) to create a correspondence between the wage reported by the employer and those reported to the tax authorities by the individual. This joint file (ATmLTO) is a much cleaner data set and is therefore used instead of the AT register. Importantly, the recorded wages are not subject to any sort of right-censoring. In addition to information about wages at the person-firm-year level, the database includes the first and the last date of the employment spell within a given year, the total number of days worked, the municipalities in which the individual lives and works, and an indicator for full-time and part-time employment. The ATmLTO data is then merged with time-varying demographic information about years of education, gender and the number of children, all from administrative registries at Statistics Norway.

We match the employer-employee with data on close to the universe of Norwegian manufacturing firms with detailed information on their international trade. The data set covers around 90 percent of the manufacturing output in Norway.⁷ The main source of the firm-level data is Statistics Norway's *Capital database*, and contains variables such as value added, output, and employment for an unbalanced panel of all mainland joint-stock firms in the Norwegian manufacturing sector, and assigns a unique identifier to each firm. The industry classification used is the NACE Rev. 2.⁸ We merge the *Capital database* with firm-level data on export values based on customs declarations. These customs data provide annual exports for all firms disaggregated by 8-digit HS product codes and destination country. Finally, we also add firm-level information on outward FDI.

2.2 Construction of the Sample and Measurement

Sample. Merging the data set for firms and employees leaves us with a panel covering the population of all mainland joint-stock manufacturing firms along with their trade, FDI and, importantly, their employees. For each worker-year combination, we assign to the worker the wage and the firm of the longest employment spell during the year. We restrict our sample to individuals who have worked for at least three months during a year. We keep only full

⁷In 2010, the manufacturing sector accounted for 26 percent of total exports, 34 percent of total goods exports and 50 percent of total mainland exports, i.e., exports excluding oil, oil-related services and ship- and platform-building. The top exporting industries within manufacturing were machinery and other equipment n.e.c.; basic metals; basic chemicals, chemical and mineral products; food products, beverages and tobacco; refined petroleum products.

⁸A list of the manufacturing industries can be found in the Appendix.

time employees in the data set. This is done to avoid biases related to possible part time wage penalties.⁹ We also restrict our sample to individuals who are between 19 and 67 years of age and to workers with at least one year of potential labor market experience. In constructing our sample we also remove outliers. To do so, we predict wages based on a simple Mincerian regression of log wages on education, experience and experience squared. See below for how the experience and education variables are constructed. We then remove observations that lie outside five times the standard error of the residual. We are left with a sample of 2,713,623 worker-firm-year observations, based on roughly 6,000 firms each year spread across 24 industries.

Wages. The outcome variable in our analysis is wages. We construct hourly wages based on information on what is referred to as cash compensation, which includes not only the basic payment but also bonuses. Non-cash forms of compensation, such as stock options and shares, are not included.

When estimating the gender wage gap we account for a host of variables to account for observable worker characteristics:

Education. We measure education using an individual's total number of years of schooling. We classify a worker as being College educated if the worker has 14 or more years of education, and non-college educated (No college) if not.

Experience. To account for work experience we use a measure of actual experience calculated from the Pension Register, and gives the actual number of years a person has been active in the labor market.¹⁰ An alternative and often used measure of labor market experience is a person's potential experience, given by the age of the worker minus the number of years of education. However, given our focus on gender differences, we choose the actual experience to be our preferred measure of labor market experience since women typically have noncontinuous labor market histories and thus potential experience may be a poor proxy for their actual experience.¹¹

Children. We include a dummy variable that is equal to one if the worker has at least one child in the year of observation, and zero otherwise. An alternative demographic characteristic would be a worker's marital status, but since cohabitation is very common in Norway also after having children, we believe a person's private life is better characterized by the Children variable.

Centrality. To account for geographical aspects of the location of firms and workers, we employ a centrality measure constructed by the Norwegian Institute for Urban and Regional Research (NIBR). All Norwegian municipalities are grouped into residential and labor market regions. These regions are then split into five categories along the centre-periphery axis, based on the size of the population, availability of amenities, number of jobs and distance to the nearest place categorized as central. The categories are: major cities, medium-sized towns, small towns, rural centres and periphery. We define Centrality variable that equals one if an

⁹It is well known in the gender wage gap literature that part-time workers have lower hourly earnings and that women are overrepresented among part-time workers, see e.g. Manning and Petrongolo (2008).

¹⁰The Pension Register contains data on incomes dating back to 1967.

¹¹Note however, that since employees earn pension rights also during parental leave, the registered number of years in the labor market includes these periods of leave.

individual is working in a municipality that belongs to one of the first two categories in the year of observation, and zero otherwise.

Exporter. In line with the literature, we classify a firm as an Exporter if the value of its total exports in the year of observation exceeds NOK 10,000 (USD 1,100). To measure the extent and complexity of exporting we use the ratio of exports to total production (Export share), the number of destinations to which a firm exports in a given year (Destinations) and the number of exported varieties (No of varieties) defined as the number of 8-digit HS product code-destination country combinations.

Business hour overlap. As a measure of the geography of exporting in a given year we use a weighted measure of business hour overlap. We compile data on time zone(s) for each export destination country. For countries spanning multiple time zones, we consider only mainland and use the average value. We calculate the the overlap in business hours between each destination country and Norway assuming the standard 9am to 5pm office hours. We then use information on firms exports across destinations and varieties. Assuming that the need for communications with a particular country increases in the number of traded products, for each firm we take the average business hour overlap across country-product combinations. For non-exporters, the overlap is equal to 8 hours.

Gender gap index. We want to control for different attitudes towards women in destination countries. We therefore use the Gender Gap Index from the World Economic Forum. It pertains to 2006, and covers 115 countries. The index measures gaps between men and women, rather than absolute levels of female outcomes,. Second, it focuses on outcomes rather than input measures. Third, it rewards countries for closing the gender gap, but neither rewards not penalizes countries where women have an advantage over men in certain outcomes. As with the Business hour overlap, we take the average of the index over all product-destination country combinations observed in a particular year.

2.3 Descriptives

In line with findings of studies from other countries, we observe that the majority of Norwegian manufacturing firms do not export. In 1996, only 34.9 percent were exporting, while in 2005, this number had risen to 38.1 percent. Figure 2 in the Appendix shows that there is also substantial heterogeneity in the share of exporting firms across Norwegian manufacturing industries.

Table 1 provides an overview of the distribution of workers between exporters and non-exporters and the average characteristics of the labor force at these two types of firms at the overall level. If anything, the two types of firms appear remarkably similar with respect to their labour force as long as we only account for observable characteristics. However, we note that while non-exporters pay an average wage to women that is 24 percent lower than what they pay to men, exporters pay women (only) 20 percent less. We see that exporters on average employ workers with slightly higher education, slightly more labor market experience and workers who

¹²This index was first compiled in 2005, but covered only 58 countries in that year.

Table 1: Exporters versus Non-Exporters

	N	on-Export	ers		Exporters			
	All	${\bf Female}$	$_{ m Male}$	All	${\bf Female}$	$_{ m Male}$		
Average wage	105.67	84.33	110.83	129.30	108.59	134.76		
Average education	12.81	12.18	12.28	12.93	12.79	12.97		
Average experience	22.33	18.83	21.61	22.93	21.02	23.04		
Average age	41.31	39.73	40.16	41.60	41.46	41.63		
Share with Children	0.9	0.88	0.89	0.81	0.91	0.91		
Average no of children	1.84	1.91	1.82	1.81	1.84	1.8		
Average Centrality	0.52			0.53				
Female share in labor force	.20			0.21				
Total share in labor force	0.19			0.81				

Notes: All numbers apart from average wages are based on the panel of worker-level data for 1996-2010. Average hourly wages (in NOK) are calculated for 2001. An exporter has by definition exports above NOK 10,000 (USD 1,100).

are slightly older, but have close to the same number of children. In line with what is observed across a wide set of countries, the exporters' employees on average have substantially higher wages. The average female share in the labor force is approximately the same for exporters and non-exporters. Within the group of manufacturing industries, the highest share of females among full-time employees can be found in wearing apparel (74.4 percent), followed by pharmaceuticals (50.6 percent) and textiles (43 percent).

3 Empirical Specification

We want to investigate whether a firm's export status matters for the GWG. Moreover, in line with the hypothesis presented in the introduction, we want to examine whether heterogeneity across exporters with respect to the extent, complexity and geography of their operations matters for the size of the GWG.

To answer these questions we adopt the workhorse tool from the existing literature and estimate a worker-level Mincer (1974) type wage regression. We augment it by adding a female dummy and its interaction with a variable indicating the firm's export status (a dummy variable) or measuring the extent, complexity or geography of exporting (a continuous variable):

$$\ln w_{ijst} = \beta_1 fem_i + \beta_2 fem_i * Exporter_{jt} + \beta_3 Exporter_{jt} + X'_{it} \gamma + \varepsilon_{ijst}. \tag{1}$$

where w_{ijst} is the hourly wage of worker i employed by firm j operating in industry s at time t, fem_i is a dummy for being female, and X_{it} is a vector of observable characteristics of worker i. It includes the variables Education, Experience, Children, and the dummy for working in a central location (Centrality). As is common practice, we also include experience squared (divided by 100). Exporter is either a dummy equal to one if firm j is an exporter at time t or a continuous

measure of its exports based on export share in sales, number of destinations, product varieties or average business hours overlap. The details on the variables and their construction were provided in Section 2.2.

A statistically significant coefficient on the female dummy captures the wage gap between males and females, assuming that the included controls adequately account for differences in worker productivity. A statistically significant coefficient on the interaction term suggests that the GWG differs between exporters and non-exporters.

However, estimating the impact of exporting on the GWG poses a set of challenges. To deal with these we estimate three different types of specifications. In the first specification, we control for industry-year fixed effects (α_{st}) to account for any systematic variations in wages across industries and thus focusing on within-industry variation. Moreover, we condition our identification on changes within a firm by adding firm fixed effects (α_j) to take into account unobservable firm heterogeneity that is likely to affect exporting as well as wages:

$$\ln w_{ijst} = \beta_1 fem_i + \beta_2 fem_i * Exporter_{jt} + \beta_3 Exporter_{jt} + X'_{it} \gamma + \alpha_{st} + \alpha_j + \varepsilon_{ijst}.$$
 (2)

In the second specification, we condition on the changes within a given worker-firm match by including job-spell fixed effects (α_{ij}) . As in the first specification, we also add industry-year fixed effects (α_{st}) :

$$\ln w_{ijst} = \beta_1 fem_i * Exporter_{it} + \beta_2 Exporter_{it} + X'_{it}\gamma + \alpha_{st} + \alpha_{ij} + \varepsilon_{ijst}. \tag{3}$$

In our final and preferred specification, we condition our identification on changes within the worker-firm match effects while also controlling for firm-year fixed effects (α_{it}). The latter captures unobserved heterogeneity that may be driving selection of firms into exporting, as well as potential unobservable confounders such as shocks to firm productivity:

$$\ln w_{ijst} = \beta_1 fem_i * Exporter_{jt} + X'_{it}\gamma + \alpha_{ij} + \alpha_{it} + u_{ijst}, \tag{4}$$

In all specifications, standard errors are clustered at the firm level to account for group correlation across workers within firms, as well as correlation within firms over time.¹³ In what follows, we present results based on all three specifications..

4 Empirical Results

4.1 Exporters and the Gender Wage Gap

The estimation results are presented in Table 2. We separately estimate the wage regressions described above splitting the sample into two subsamples: college educated and non-college educated workers. We believe that our hypothesis about exporters demanding more flexibility

¹³All of our results are robust to clustering at the level of individual workers.

and commitment applies particularly to college educated workers. Therefore, we expect to see a differential effect between exporters and non-exporters in terms of the GWG for this sub-sample.. For this reason, we first present the results for the college educated and only later comment on the results for the non-college educated.

The results in Column (1) are based on the estimation of specification (1). They show that after controlling for education, experience and children as well as centrality of the workplace location, college educated women earn roughly 23 percent less than men, but that this gap is slightly lower at exporting firms (around 19 percent). Thus working at an exporter seems to have a substantial impact on closing the gender wage gap. The estimated coefficients are statistically significant at the one percent level. The coefficients on the other variables have the expected signs, and their magnitudes are in line with the existing literature. Both experience and education are positively correlated with wages, while having children has a negative effect. 14 We also observe that workers working for more centrally located firms tend to earn more. Our first set of results reveal differences in the gender wage gap between exporters and non-exporters. But the women working for exporters may be different from women working for non-exporters in ways that the econometrician cannot observe. The lower gender wage gap observed in exporting firms can simply be an artifact that can be explained by exporters employing different women than non-exporters. Such differences in firms' labor force could be explained by labor market sorting, where women with greater abilities and ambition seek to be employed by exporters. They can also arise if exporters are better at screening and thus employ women with higher productivity than non-exporters.

To account for such unobservables we estimate specification (2), i.e. we control for unobservable characteristics that determine the match between workers and firms by including spell fixed effects. A spell is defined as a unique worker-firm combination. This estimation strategy uses variation coming only from firms changing their export status, holding all worker- and match-specific time-invariant effects constant. In this way, we circumvent the problem of endogenous mobility that occurs when relying on worker fixed effects, and we are are able to account for firm-specific unobservable characteristics at the same time. As these fixed effects will account for all time-invariant heterogeneity of the workers, we are not able to identify the overall gender wage gap using this methodology.¹⁵ As before, we also include sector-year fixed effects.

The results, presented in Column (2), tell a very different story than the ones presented in Column (1). The coefficient on our interaction variable Female * Exporter is statistically significant at the one percent level, but now it bears a negative sign for college graduates. In other words, when we account for unobservable heterogeneity specific to the worker-firm match, we find that the GWG is larger in exporting firms than in non-exporters. This suggests that the observable characteristics that we included in our previous regressions did not capture all the relevant aspects of a worker's productivity and the quality of the worker-firm match and that

¹⁴It is worth recalling at this point that *Experience* squared enters divided by 100.

 $^{^{15}\}mathrm{No}$ workers in our sample change gender.

Table 2: GWG: Controlling for Observables and Unobservables

	College (1)	College (2)	College (3)	No College (4)	No College (5)	No College (6)
Female*Exporter	.040***	032***	030**	.061***	008	004
	(.012)	(.010)	(.014)	(.006)	(.005)	(.005)
Female	257***			265***		
	(.012)			(.004)		
Exporter	007	.011*		007	.006*	
	(.009)	(.006)		(.003)	(.004)	
Education	.061***	.063***	.061***	.039***	.075***	.075***
	(.002)	(.008)	(.008)	(.001)	(.003)	(.003)
Experience	.038***	.014	$.021^{*}$.032***	.041***	.037***
	(.001)	(.011)	(.012)	(.000)	(.010)	(.010)
Experience squared	053***	066***	067***	055***	051***	051***
	(.005)	(.002)	(.002)	(.002)	(.001)	(.001)
$\operatorname{Children}$.017***	035***	033***	035***	083***	084***
	(.005)	(.009)	(.010)	(.002)	(.006)	(.006)
Centrality	.101***	.010	.011	$.051^{***}$	003	006
	(.018)	(.007)	(.010)	(.010)	(.008)	(.005)
Industry-Year FE	Yes	Yes	No	Yes	Yes	No
Firm FE	Yes	No	No	Yes	No	No
Spell FE	No	Yes	Yes	No	Yes	Yes
Firm-Year FE	No	No	Yes	No	No	Yes
Obs	553,611	553,611	553,611	$2,\!160,\!012$	2,160,012	2,160,012
R-squared	.50	.81	.89	.40	.70	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. ***= p-val<.01, ** = p-val<.05, * = p-val<.1

exporters do indeed hire more "able" workers, in terms of unobservable characteristics. Once these unobservable characteristics are taken into account, there is a wage penalty for college educated women working in exporting firms, equal to about 3 percentage points. We also find an exporter premium of about 1 percentage point, which is in line with the other studies based on linked employer-employee data.

In our final and preferred specification (3), we replace industry-year fixed effects with firm-year fixed effects. In this way, we control for firm-specific time-varying shocks that may be responsible for self-selection of firms into exporting. As visible in Column (3), doing so has little impact on the coefficient of interest, which remains negative and statistically significant. Again we find a female wage penalty of about 3 percentage points among workers with college education.

Taking all these results together, we conclude that although college educated women working for exporting firms earn higher wages than college educated women employed by non-exporters, they are underpaid relative to college educated men given their unobservable characteristics.

Workers without College Education Are Not Affected Let us now turn to the estimates conducted on the subsample of workers without college education, presented in the right-hand-side panel of Table 2. As before, when we control just for observable worker characteristics, we find a lower GWG in exporting firms. However, once spell fixed effects are included in the regressions, this effect ceases to be statistically significant while changing sign. Hence, accounting both observables and unobservables, we do not find evidence of a bigger GWG at exporting relative to non exporting firms with respect to non-college educated workers. This is not surprising as our hypothesis about exporters requiring their workers to work particular hours to communicate with customers in foreign markets is unlikely to apply to production workers.

4.2 Exporter Heterogeneity and the Gender Wage Gap

So far our analysis has focused on the differential gender wage gaps between exporters and non-exporters. However, exporters themselves are not a homogenous group. Exporters vary according to the extent to which they rely on exports rather than domestic sales. They also differ with respect to the complexity of their operations which is reflected in the number of export destinations and the number of product varieties exported (a product variety is defined as the product-destination combination). As visible in Table 3, there is quite a lot of heterogeneity among exporters. The median exporter exports 4 percent of its sales, serves three export destinations and supplies six varieties. There are, however, exporters serving just one foreign market and those serving 146 markets. Similarly, the sample includes exporters supplying one variety as well as those with 4,126 varieties. There is also quite a lot of variation in the business hour overlap, which ranges from 8 hours to zero, with 7.67 being the mean value and quite many values found inbetween (see Figure 1).

Table 3: Exporter descriptives

	1			
	Mean	Median	Min	Max
Export share	0.18	0.04	< 0.01	1
Business hour overlap	7.01	7.67	0	8
No of destinations	7.31	3	1	146
No of varieties	24.88	6	1	$4,\!126$

Notes: All numbers are based on the panel of firm-level data for 1996-2010. An exporter has by definition exports above NOK 10,000 (USD 1,100). A variety is defined as product-destination combination. <0.01 denotes export share below 1 percent.

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Figure 1: Business hours overlap

4.2.1 Size and Complexity of Export Operations and the Gender Wage Gap

Exporting a larger share of output or supplying a larger number of destinations will typically require more interactions with foreign customers. The same will most likely be true of exporting a larger number of products, because there is a positive relationship between the number of products supplied to a market and the number of buyers (see e.g. Carballo et al. (2013)). Therefore, we would expect bigger and more complex export operations to be associated with a larger GWG.

We now set out to investigate the role of heterogeneity among exporters in the GWG. We estimate our preferred and most stringent specification on the subsample of firms that are exporters in the year of observation allowing the GWG to differ with the continuous measures of exports: (i) the share of output exported; (ii) the number of export destinations; (iii) the number of exported varieties (where a variety is defined as a product-country combination).

Table 4: GWG: Accounting for heterogeneity among exporters

	College	College	College	No College	No College	No College
	(1)	(2)	(3)	(4)	(5)	(6)
Female*Export share	014**			011		
	(.007)			(.011)		
Female * Destinations		013**			007*	
		(.006)			(.003)	
Female*Product varieties			009**			006**
			(.004)			(.002)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	454,063	494,601	494,601	$1,\!620,\!597$	1,712,387	1,712,387
R-squared	.88	.88	.88	.79	.80	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Only workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls include education, experience, experience squared, children and centrality. ***= p-val<.01, ** = p-val<.05, * = p-val<.1

The results in Table 4 are supportive of our hypothesis. We find that the GWG increases as each of the three measures of export intensity increases. The estimated coefficients of interest are statistically significant at the five percent level when the analysis focuses on college educated workers. Interestingly, we also find similar, albeit smaller in magnitude and less robust, evidence for workers without college education. Focusing on the college educated workers, an increase in the export share of 10 percentage points is associated with the GWG going up by 13 percentage points. Adding four additional export markets leads to the GWG going up by five percentage points. Finally, adding ten additional export varieties means an 8 percentage point increase in the earning differential between men and women. Hence, our results suggest that the magnitude and the complexity of the export operations matters for the GWG and the bigger and more complex operations are associated with a greater the GWG.

4.2.2 The Geography of Exporting and the Gender Wage Gap

If our hypothesis that the need for working particular hours translates into a higher GWG holds, then it is not just the number of markets that should matter but also their location. In particular, as argued earlier, what should matter is the overlap in business hours. With a smaller overlap the chances are that the interactions with customers will need to take place either before or after the standard office hours. Taking early morning or late night phone calls makes it more difficult for employees to keep the balance between their personal life and professional commitments. Moreover, a larger time difference means longer travel time required to visit the market in question and thus the need to be away from home for a longer time period. If women

Table 5: GWG: Exporting and business hours overlap, I

	All Workers							
	College	College	College	No College	No College	No College		
	(1)	(2)	(2)	(3)	(4)	(4)		
Female* ln Business hours overlap	.028*		.033**	.001		.001		
	(.015)		(.015)	(.009)		(.009)		
Female* ln Gender index		014	082		.006	.012		
		(.099)	(.102)		(.048)	(.052)		
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes		
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes		
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes		
Obs	493,736	494,601	493,177	1,708,945	1,709,121	1,705,848		
R-squared	.88	.88	.88	.80	.80	.80		

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Only workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls include education, experience, experience squared, children and centrality. ***= p-val<.01, ** = p-val<.05, * = p-val<.1

shy away from such duties, or are perceived as being less able or less willing to accommodate them, the GWG should increase as the business hours overlap goes down.

To test the hypothesis that where you export matters for the GWG, we generate a measure of business hour overlap based on the destination markets of the firm. As argued above, the need for communication with customers is likely to increase with the number of exported varieties. Therefore, we specify the overlap in business hours for each product-destination country combination and take an average over all varieties the firm is exporting.

Table 5 presents the estimation results from a model where we interact the log of this business overlap variable with the female dummy. The estimated coefficient on this interaction term is positive and statistically significant for college educated workers (column (1)) but – in line with expectations – not for workers with less education (column (4)). A four hour decrease in the business hour overlap is associated with the GWG increasing by 11 percentage points for college educated women.

One may be concerned that the business hour overlap is capturing societal attitudes, such as the position of women in the society of the export destination, instead of just purely reflecting the time zones where a firm's customers are located. It is possible that it is harder for female employees of Norwegian exporters to do business in countries where women's rights are restricted or women have a lower social status. Countries where this is typically known to be the case are typically far away from Norway and will thus have small business hour overlap. If so, our results could simply be reflecting lower effectiveness of female employees due to the lack of willingness among the firms' foreign partners to do business with women, rather than the need for working outside the standard office hours due to time zone differences.

Table 6: GWG: Exporting and business hours overlap, II

		Workers under 45 years of age								
	College (1)	College (2)	College (3)	No College (4)	No College (5)	No College (6)				
Female* ln Business hours overlap	.042**		.049**	.005		036				
Female* ln Gender index	(.020)	027	(.021) 129	(.012)	030	(.080) 036				
		(.124)	(.133)		(.074)	(.013)				
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes				
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes				
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes				
Obs	327,973	328,142	327,559	$1,\!028,\!554$	1,028,674	1,026,584				
R-squared	.88	.88	.88	.81	.81	.81				

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Only workers employed by exporting firms are included. Standard errors in parentheses clustered by firm. Worker controls include education, experience, experience squared, children and centrality. ***= p-val<.01, ** = p-val<.05, * = p-val<.1

To address this concern we control for attitudes towards women and gender equality at the export destinations of each Norwegian exporter. To to so we use the Gender Gap Index which is compiled by the World Economic Forum and pertains to year 2006 (see Section 2.2 for more information about the index). The value of the index relevant to each exporter is calculated in a way analogous to the business hour overlap. In Column (2), we first show that there is no systematic relationship between the Gender Gap Index relevant to each exporting firm and the GWG in that firm. In other words, we show that an interaction term between the female dummy and the index is not statistically significant in either subsample. Then in Column (3) we show that the relationship between the business hours overlap and the GWG is robust to allowing for the GWG to differ with the gender inequality index in the export markets. Controlling for the Gender Gap Index actually increases the significance level of the business hour overlap variable.¹⁶

Our results suggest that where you export matters. The smaller the overlap in business hours the greater the wage gap. Regardless of whether this can be related to women being less flexible in terms of availability around the clock or whether this is about employers perception of this being the case, we may expect younger women (who are likely to have young children) to be more affected than older women. As fertility decisions are not exogenous, we restrict our attention to workers of both genders under 45 years of age and re-estimate the specifications from the previous table on this subsample. Again we focus on the college educated workers, and find that, as expected, a lower business hour overlap is associated with a higher GWG differential in

 $^{^{16}}$ Not logging the business hour overlap or the Gender Gap Index leads to the same conclusions. See Appendix Table 9.

this subsample than in the population of workers (see Table 6). The estimated magnitudes are economically meaningful. A two-hour decline in the business hour overlap increases the GWG by 6.7 percentage points in the full sample and by 10 percentage points in the subsample of workers under 45.

4.3 Addressing an Alternative Explanation

Finally, we turn to a possible alternative interpretation for our baseline findings of a higher GWG among exporting firms. Becker's (1957) theory of taste-based discrimination predicts that only very profitable firms can afford to discriminate. As exporters are more profitable than other firms, they may be *better* positioned to engage in costly discrimination against women. In other words, Becker's theory would lead to the same prediction of a higher gender wage gap among exporters, but its interpretation would be different.

To address this alternative explanation we augment our model by allowing the GWG to differ with firm profitability, measured as the log of profits divided by the operating income.¹⁷ We also allow it to differ with firm size, measured as logged employment, and which in the Melitz (2003) framework is positively correlated with profitability. Finally, we account for its multinational status, as is known to be highly correlated with profitability.

As evident from Table 7, doing so does not affect our results. As before, we find that a lower business hour overlap is associated with a larger GWG for college educated workers. The magnitude of the effect and its significance level is the same as those found in Table 5. This is true regardless of whether we control for the Gender Gap Index. As for the additional controls, they do not appear to be correlated with the GWG, with the exception of firm profitability. The results suggest that the GWG is higher in more profitable firms, which is in line with Becker's theory.

¹⁷The *Profitability* variable is missing for approximately 22 percent of the sample. Instead of dropping these observations, we add a dummy for missing *Profitability* in all the specifications where *Profitability* is included.

Table 7: GWG: Alternative explanations

	College (1)	College (2)	College (2)	No College (3)	No College (4)	No College (4)
Female*ln Business hours overlap	.028*		.034**	.001		.001
	(.015)		(.015)	(.009)		(.009)
Female* ln Gender index		014	082		.006	.012
		(.099)	(.102)		(.048)	(.053)
Female*Profitability	000	000	000	002**	002**	002**
	(.001)	(.001)	(.001)	(.005)	(.001)	(.001)
Female*MNC	002	002	002	.002	.002	.001
	(.008)	(.008)	(.008)	(.006)	(.006)	(.006)
Female*ln Size	004	003	004	008	008	008
	(.008)	(.008)	(.008)	(.005)	(.005)	(.005)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	493,736	494,016	493,177	1,708,945	1,709,121	1,705,848
R-squared	.88	.88	.88	.80	.80	.80

Notes: Dependent variable is log wage. MNC is a dummy that equals one if the firm is registered as having a positive ownership share in a firm located in a foreign country in the year of observation, and zero otherwise. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. Worker controls include education, experience, experience squared, children and centrality. ***= p-val<.01, *** = p-val<.05, * = p-val<.1

5 Conclusions

We hypothesize that exporters, by virtue of doing business with customers located in remote markets and different time zones, disproportionately reward flexibility and availability round the clock of their employees. If women are less able to accommodate such demands, or are perceived as such, exporting firms will exhibit a larger GWG. Due to the different nature of work carried out by college versus non-college educated workers, we expect this to be true in particular for white collar workers.

We test our hypothesis using matched employer-employee data from the Norwegian manufacturing sector for the 1996-2010 period. The richness of the data allows us to control for unobservable employer-employee match heterogeneity as well as for unobservable firm-year shocks. As expected, we find that for college educated workers exporting firms have on average a 3 percentage points higher GWG than non-exporters. Further, we exploit the fact that there is substantial heterogeneity among exporters with respect to the magnitude and complexity of exports. We show that the GWG increases with the share of output exported, the number of markets served and the number of export varieties. Exporters also differ with respect to the

geography of the markets that they are supplying. A priori we would expect the handling of far away destinations in different time zones to be more demanding. To investigate this we account for the average overlap in business hours a firm faces across export destinations. This allows us to test and confirm our hypothesis that the GWG is related to the need for communications outside the standard business hours in the home country. This result proves to be robust to controlling fordifferences in gender inequality and female rights across destination countries. Our hypothesis is further supported by sub-sample results showing that the magnitude of the effect is larger for workers in their child bearing years.

Our results demonstrate another, perhaps somewhat unexpected, channel through which firms' internationalization affects income inequality. Building on Goldin (2014) one may thus argue that the last chapter of gender convergence is not only about certain sectors but also about certain firms and the need for making sure that the gains from trade are distributed in such a way that they reduce rather than enhance inequalities.

References

- Amiti, M. and D. Davis (2012). Trade, firms, and wages: Theory and evidence. Review of Economic Studies 79, 1–36.
- Barth, E., S. P. Kerr, and C. Olivetti (2017, May). The dynamics of gender earnings differentials: Evidence from establishment data. NBER Working Papers 23381, NBER Working Paper 23381.
- Becker, G. S. (1957). The Economics of Discrimination. Chicago: University of Chicago Press.
- Ben Yahmed, S. (2013). Gender wage gaps across skills and trade openness. Working paper.
- Bernard, A. B., J. B. Jensen, and R. Z. Lawrence (1995). Exporters, jobs, and wages in US manufacturing: 1976-1987. *Brookings Papers on Economic Activity. Microeconomics* 1995, 67-119.
- Black, S. E. and E. Brainerd (2004). Importing equality? The impact of globalization on gender discrimination. *Industrial and Labor Relations Review* 57(4), 540–559.
- Blair-Loy, M. and J. A. Jacobs (2003). Globalization, work hours, and the care deficit among stockbrokers. Gender & Society 17(2), 230–249.
- Blau, F. D. and L. M. Kahn (2000). Gender differences in pay. *Journal of Economic Perspectives* 14(4), 75–99.
- Carballo, J., G. I. P. Ottaviano, and C. V. Martineus (2013). The buyer margins of firms' exports. CEPR Discussion paper no 9584.
- Card, D., A. R. Cardoso, J. Heining, and P. Kline (2016). Firms and labor market inequality: Evidence and some theory. Technical Report 22850, NBER Working Paper no 22850.
- Gislason, I. (2007). Parental leave in Iceland: Bringing the fathers in. Technical report, Centre for Gender Equality, Ministry of Social Affairs.
- Goldberg, P. and N. Pavcnik (2005). Trade, wages, and the political economy of trade protection: Evidence from the colombian trade reforms. *Journal of International Economics* 65, 75–105.
- Goldberg, P. and N. Pavcnik (2007). Distributional effects of globalization in developing countries. *Journal of Economic Literature* 45(1), 39–82.
- Goldin, C. (2014). A grand gender convergence: Its last chapter. American Economic Review 104 (4), 1091–1119.
- Hirsch, B., T. Schank, and C. Schnabel (2010). Differences in labor supply to monopsonistic firms and the gender pay gap: An empirical analysis using linked employer–employee data from germany. *Journal of Labor Economics* 28(2), 291–330.

- Irarrazabal, A., A. Moxnes, and K. H. Ulltveit-Moe (2013). Heterogeneous firms or heterogeneous workers? Implications for the exporter premium and the impact of labor reallocation on productivity. *Review of Economics and Statistics* 95(3), 839–849.
- Juhn, C., G. Ujhelyi, and C. Villegas-Sanchez (2014). Men, women and machines: How trade impacts gender inequality. *Journal of Development Economics* 106, 179–193.
- Krishna, P., P. Jennifer, and M. Z. Senses (2014). Wage effects of trade reform with endogenous worker mobility. *Journal of International Economics* 93, 239–252.
- Kvande, E. (2009). Work-life balance for fathers in globalized knowledge work. some insights from the norwegian context. Gender, Work and Organization 16(1), 58–72.
- Manning, A. and B. Petrongolo (2008). The part-time pay penalty for women in Britain. *The Economic Journal* 118, F28–F51.
- Melitz, M. J. (2003). The Impact of Trade on Intra-Industry Reallocations and Aggregate Industry Productivity. *Econometrica* 71(6), 1695–1725.
- Mincer, J. (1974). Schooling, Experience, and Earnings. Columbia University Press.
- OECD (2012). Closing the Gender Gap: Act Now. OECD Publishing.
- Pavcnik, N., A. Blom, P. Goldberg, and N. Schady (2004). Trade liberalization and industry wage structure: Evidence from brazil. World Bank Economic Review 18(3), 319–344.
- Revenga, A. (1997). Employment and wage effects of trade liberalization: The case of mexican manufacturing. *Journal of Labor Economics* 15, S20–S43.
- Schank, T., C. Schnabel, and J. Wagner (2007). Do exporters really pay higher wages? First evidence from German linked employer–employee data. *Journal of International Economics* 72(1), 52–74.
- Trefler, D. (2004). The long and the short of the canada-us free trade agreement. American Economic Review 94, 870–895.

A Appendix

A.1 Manufacturing Industries

Table 8: NACE Rev. 2

	Table 6. NACE Rev. 2
NACE code	Industry name
10	Manufacture of food products
11	Manufacture of beverages
12	Manufacture of tobacco products
13	Manufacture of textiles
14	Manufacture of wearing apparel
15	Manufacture of leather and related products
16	Manufacture of wood and products of wood and cork, except furniture;
	manufacture of articles of straw and plaiting materials, except furniture
17	Manufacture of paper and paper products
18	Printing and reproduction of recorded media
19	Manufacture of coke and refined petroleum products
20	Manufacture of chemicals and chemical products
21	Manufacture of basic pharmaceutical products and pharmaceutical preparations
22	Manufacture of rubber and plastic products
23	Manufacture of non-metallic mineral products
24	Manufacture of basic metals
25	Manufacture of fabricated metal products, except machinery and equipment
26	Manufacture of computer, electronic and optical products
27	Manufacture of electrical equipment
28	Manufacture of machinery and equipment n.e.c.
29	Manufacture of motor vehicles, trailers and semi-trailers
30	Manufacture of other transport equipment
31	Manufacture of furniture
32	Other manufacturing
33	Repair and installation of machinery and equipment

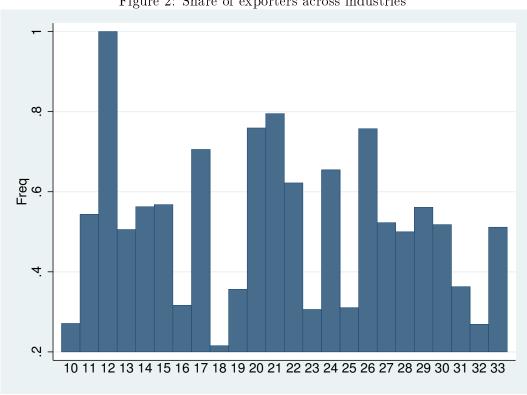


Figure 2: Share of exporters across industries

Table 9: GWG: Exporting and business hours overlap, Alternative specification

1	<u> </u>			All Workers	1	
	College	College	College	No College	No College	No College
	(1)	(2)	(2)	(3)	(4)	(4)
Female*Business hours overlap	.005*		.006**	.000		.000
	(.003)		(.003)	(.001)		(.002)
Female* Gender index		023	132		.006	.009
		(.136)	(.141)		(.065)	(.072)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	494,575	494,601	494,601	1,712,218	1,709,121	1,705,848
R-squared	.88	.88	.88	.80	.80	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. Worker controls include education, experience, experience squared, children and centrality. ***= p-val<.01, ** = p-val<.05, * = p-val<.1